

# Do age-specific survival patterns of wild boar fit current evolutionary theories of senescence?

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### 1 Do age-specific survival patterns of wild boar fit current evolutionary

### 2 theories of senescence?

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- 18
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Actuarial senescence is an inevitable outcome of evolution in age-structured populations. In growing populations, the progressive decline of Hamiltonian forces of selection with age leads to decreasing survival. As actuarial senescence is over-compensated by a high fertility, actuarial senescence should be more intense in species with high reproductive effort, a theoretical prediction that has not been yet explicitly tested across species. Wild boar (Sus scrofa) has an unusual life-history strategy among large mammals by associating both early and high reproductive effort with potentially long lifespan. Therefore, wild boar should show stronger actuarial senescence than similar-sized related mammals. Moreover, being polygynous and much larger than females, males should display higher senescence rates than females. From a long-term monitoring of 1,783 individuals, we tested these predictions. We provided clear evidence of actuarial senescence in both sexes. Wild boar females had earlier but not stronger actuarial senescence than similar-sized ungulates and both sexes displayed similar senescence rates. Our study indicates that the timing of senescence, not the tempo, is associated with the magnitude of fertility in ungulates. This demonstrates the importance of including the timing of senescence in addition to its rate to understand variation in senescence patterns in wild populations. 

49 Actuarial senescence (hereafter senescence) defined as the decline of survival with 50 increasing age, has been repeatedly documented in wild species of vertebrates (see Nussey et 51 al. 2013 for a review). However, empirical evidence has increased only recently, 52 thanks to the development of long-term individual monitoring of known-aged animals 53 (Clutton-Brock and Sheldon 2010). William D. Hamilton already demonstrated nearly fifty-54 years ago (Hamilton 1966) that senescence was an inevitable outcome of evolution in age-55 structured populations. More specifically, based on the decline of the forces of natural 56 selection with age in growing age-structured populations (Medawar 1952), Hamilton predicted that an increase in fertility should be selected at the costs of a survival decrease at 57 58 old ages. In other words, high fertility should be associated with high rates of senescence 59 (Hamilton 1966). Until now, intraspecific studies have provided good evidence for a 60 relationship between early allocation to reproduction and fitness-related traits, such as 61 survival (Hammers et al. 2013) and components of reproductive success (Nussey et al. 2006) 62 in late life, in line with the disposable soma theory of ageing (Kirkwood 1977; Kirkwood and 63 Austad 2000). On the other hand, while some recent comparative studies have revealed an 64 association between intensity of actuarial senescence and pace of life across species within 65 and among classes of vertebrates (Jones et al. 2008, Péron et al. 2010, Lemaître and Gaillard 66 2013), no study has tested yet the exact Hamilton's prediction of a direct link between 67 fecundity and intensity of actuarial senescence across species.

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Moreover, while empirical evidence that females live much longer than males in most mammals studied so far is compelling (Carey and Judge 2000), whether and how age-specific trajectories differ between sexes remains an open question (Nussey et al. 2013, Maklakov and Lummaa 2013), although the theoretical basis for such sex differences has been established for a long time (Williams 1957). Hamiltonian forces of selection again provide the

74 explanation for more intense senescence in males when these latter suffer from higher 75 mortality during the prime-age stage. Starvation and antagonist contests during the mating season are likely to cause increased mortality rates in males compared to females in 76 77 polygynous species (Bonduriansky et al. 2008), especially when environmental conditions are 78 harsh (Toïgo and Gaillard 2003). Thus, under current life history theory, males from highly 79 polygynous and dimorphic species should exhibit a 'live fast, die young' life history strategy 80 (Vinogradov 1998, Bonduriansky et al. 2008, Festa-Bianchet 2013) and should therefore have 81 higher senescence rate than females.

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83 Using a 17 year-long monitoring of an Italian population of wild boar (Sus scrofa), we 84 filled these gaps by investigating the pattern of actuarial senescence for this large mammal 85 characterized by a unusual association between high fertility and high sexual size dimorphism 86 (Carranza 1996). Indeed, compared to similar-sized related species, wild boar females give 87 birth for the first time earlier (1 vs. 2-3 years of age, Servanty et al. 2009), at a lower body 88 mass (40% vs. 80% of adult body mass, Servanty et al. 2009), and produce larger litters (>3 89 vs. 1 or 2, Gamelon et al. 2013). By allocating early and heavily to reproduction, wild boar 90 females have an unusual life-history strategy among large mammals (Focardi et al. 2008). 91 Therefore, in accordance with Hamilton's theory, wild boar females should show faster 92 senescence than reported for similar-sized ungulates. Moreover, as wild boar males are 93 markedly heavier than females (90 vs. 70 kg for males and females, respectively, Gaillard et 94 al. 1992) and are highly polygynous (Delgado et al. 2008), they should suffer from more 95 intense senescence than females.

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## 99 Materials and methods

### 100 **S** 101

### STUDY SITE AND DATA COLLECTION

We studied a weakly hunted and growing wild boar population in the fenced 6,000 ha preserve of Castelporziano (41°44'N, 12°24'E) in Italy (Focardi et al. 2008). From 1995 to 2012, a total of 1,783 juvenile (< 1 year of age) and yearling (1 to 2 years old) wild boars were marked with numbered and colored ear tags and released after handling. Between July and September, wild boars were artificially fed three to six times per week on 80-88 feeding sites (Franzetti et al. 2012) to allow counts and identification of marked animals from year to year.

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# ESTIMATING AGE-SPECIFIC SURVIVAL, RATE AND ONSET OF SENESCENCE IN MALE AND FEMALE WILD BOAR

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We estimated natural survival rates of marked individuals using Capture-MarkRecapture analyses (Lebreton et al. 1992) of both recaptures and resightings. Animals
removed under population control plans were right-censored to allow estimating natural
survival. As the preserve was fenced, no dispersal occurred so that true survival was
estimated.

118 We first tested the goodness-of-fit (GOF) (Pradel et al. 2005) of the survival models 119 using U-CARE (Choquet et al. 2009a). We accounted for heterogeneity in the detection 120 process of juveniles (see Results) by modelling two different recapture probabilities according 121 to whether an individual was previously detected (trap-dependence: Gimenez et al. 2003; 122 Pradel and Sanz-Aguilar 2012). As capture effort varied among years (Focardi et al. 2008), we 123 included time-dependence in recapture probabilities. For survival probabilities, we fitted both 124 a model without any effect of age and the full age-dependent model. From the latter, the rate of senescence was measured as the slope of the regression of log-transformed age-specific 125

126 mortality rates on age to compare with senescence rates previously reported in ungulates

127 (Gaillard et al. 2003). To avoid confounding effects of growth or reproduction, we analyzed

128 mortality rates from 2 years of age onward, when growth has been completed.

129 To estimate the age at the onset of senescence, we constrained survival to vary linearly 130 on the logit scale over different ranges of ages (ages 1, 2, 3, 4, 5, or 6; see Nichols et al. 1997 131 for a similar approach and table 1). The age at onset of senescence was defined as the age 132 when survival started to decline as assessed from the selected model. We thus fitted age-133 specific survival models while testing for between-sex differences in age-specific survival and rate of senescence. We used the Akaike's Information Criterion for model selection (Burnham 134 135 and Anderson 2002), selecting the model with the lowest AIC. All analyses were implemented 136 in E-SURGE (Choquet et al. 2009b).

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# 138 SURVIVAL PATTERNS OF WILD BOAR COMPARED TO SIMILAR-SIZED 139 UNGULATES 140

141 To test whether wild boar females show faster and/or earlier senescence than reported 142 for similar-sized ungulates, we accounted for the confounding effect of body mass (Calder 143 1984). Indeed, it is well-established that the pace of life decreases with increasing mass 144 (Gaillard et al. 2005), leading large species to exhibit later onset and smaller rate of 145 senescence than small ones (Jones et al. 2008, Péron et al. 2010, Ricklefs 2010). We thus 146 assessed the relationship between senescence rate (on a log-scale) and female body mass (on a 147 log-scale) as well as the relationship between onset of senescence (on a log-scale) and female 148 body mass (on a log-scale) of similar-sized ungulates. Females of 9 intensively monitored 149 populations belonging to 7 different ungulate species were included in the analysis, namely 150 pronghorn (Antilocapra americana), Alpine ibex (Capra ibex), roe deer (Capreolus 151 capreolus), mountain goat (Oreamnos americanus), bighorn sheep (Ovis canadensis), red deer

(*Cervus elaphus*) and Pyrenean chamois (*Rupicapra pyrenaica*). Senescence rates were
extracted from Gaillard et al. (2003)'s study and onsets of senescence from Péron et al.
(2010)'s study.

155 Patterns of senescence across species might be not independent as a result of common 156 ancestry (Harvey and Pagel 1991). We thus controlled for this non-independence by using 157 Phylogenetic Generalized Least-Squares models (PGLS) that provide an index of the 158 phylogenetic signal called  $\lambda$ , varying in most of the cases between 0 (corresponding to the 159 complete absence of phylogenetic structure) and 1 (observed pattern predicted by the 160 phylogeny; see Freckleton et al. 2002). This parameter  $\lambda$  was then introduced in the model to 161 control for the phylogenetic effect (Freckleton et al. 2002). We used the phylogeny derived 162 from the phylogenetic supertree of mammals, which provides information on both topology 163 and branch length (Bininda-Emonds et al. 2007). However, in the models linking rate and 164 onset of senescence (on a log-scale) to body mass (on a log-scale) in similar-sized ungulates, 165 the phylogenetic signal ( $\lambda$ ) did not differ from 0 ( $\lambda_{rate} < 0.0001$  and  $\lambda_{onset} < 0.0001$ ), meaning 166 that there was no detectable phylogenetic signal in both rate and onset of senescence. 167 Therefore, we conducted inter-specific analyses linking senescence rate and onset of 168 senescence to female body mass, using standard linear models (i.e. without any correction for 169 phylogeny) to avoid biased estimates (Revell 2010). From these linear models, expected rate 170 and onset of senescence of wild boar females were estimated and compared with rates and 171 onsets found in populations of similar-sized ungulates. All analyses were performed with the 172 R software (R Development Core Team 2012).

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### 176 **Results**

# 178 ESTIMATING AGE-SPECIFIC SURVIVAL, RATE AND ONSET OF SENESCENCE 179 IN MALE AND FEMALE WILD BOAR 180

181 By only including individuals first captured as juveniles, the GOF test indicated some lack of fit for both sexes (global test: for male juveniles:  $\chi^2$  (61) = 102.743, P < 0.0001; for 182 female juveniles:  $\chi^2$  (72) = 128.411, P < 0.0001). More precisely, trap-dependence was 183 184 detected for both sexes and accounted for in the model. On the contrary, the GOF test indicated no lack of fit for individuals first captured as yearlings (global test: for males:  $\chi^2$ 185 (27) = 31.273, P = 0.260; for females:  $\chi^2$  (33) = 27.575, P = 0.734). Tested models thus 186 satisfactorily fitted data when trap-dependence was included in capture probabilities of 187 juveniles. 188

The best model included marked between-sex differences at age 2 (Fig. 1), a linear relationship between survival and age from 3 years old onwards, and year-dependent capture probabilities (table 1, appendix S1). Males and females had same survival at age 1 ( $\Phi_{age1}$ = 0.68 [0.64; 0.72]), but males survived a bit less than females at 2 years old ( $\Phi_{age2 male} = 0.57$ [0.50; 0.64],  $\Phi_{age2 female} = 0.85$  [0.77; 0.91]). From 3 years old onwards, males consistently had lower survival than females (Fig. 1) but both sexes had similar senescence rates (slope on the logit scale = -0.15 [-0.23; -0.08]).

196 To compare senescence rates between wild boar and similar-sized ungulates, 197 senescence rates were re-estimated using the linear regression between mortality rates (on a 198 log scale) and age. These rates exactly correspond to the Gompertz rates (i.e. the exponential 199 rate of increasing mortality with age) and slightly differ from the rates estimated directly from 200 analyses provided above (which do not exactly correspond to Gompertz rates, Loison et al. 201 1999). After removing survival rates at 10 and 12 years old for males and females, 202 respectively, estimated on the bounds of the interval [0, 1] (see Fig. 1), we found evidence of actuarial senescence in both sexes (rates of 0.11 (SE: 0.04) for males (Wald test: 9.7, df = 1, p-203

value = 0.002) and 0.14 (SE: 0.04) for females (Wald test: 16.0, df = 1, p-value < 0.0001),</li>
Fig. 2).

# 207 SURVIVAL PATTERNS OF WILD BOAR COMPARED TO SIMILAR-SIZED 208 UNGULATES 209

As expected, across ungulate species, the senescence rate tended to decrease (slope: -0.11, SE: 0.18) and the onset of senescence to be delayed (slope: 0.17, SE: 0.12) when body mass increased (Fig. 3). From the relationship between senescence rate and body mass (on a log-log scale) performed using 9 populations belonging to 7 different species, the expected rate of senescence for wild boar females was similar to the expectation (expected value of 0.16 vs. observed value of 0.14, Fig. 3a). On the other hand, the onset of senescence was much earlier than expected. From the relationship between the onset of senescence and body mass (on a log-log scale), the expected value was 8.56 years, whereas it was estimated to only 3 years in our analysis (Fig. 3b). Wild boar females had thus markedly earlier, but not stronger senescence (Fig. 3) for their body size compared to other ungulates. 

### **Discussion**

231 Our results provide clear evidence for actuarial senescence in male and female wild 232 boar and add to the spate of studies that have recently found senescence in free-living 233 vertebrates (Nussey et al. 2013). More importantly, using a pioneering inter-specific test of the 234 Hamilton's prediction of a link between fecundity and intensity of actuarial senescence, our 235 findings show that wild boar females have earlier, but not stronger senescence than related 236 species once the confounding effects of body size has been accounted for. Indeed, our 237 findings emphasize that the high reproductive output of wild boars (with litter size up to 14, 238 Servanty et al. 2007) is associated with earlier, but not stronger decline in survival with age. 239 As Hamilton's theory was strictly focused on senescence rates, our study did not provide a 240 direct support for it. However, we show that actuarial senescence occurred much earlier in 241 wild boar females than in other ungulates. High and early fertility in wild boar females during 242 the prime-age stage might advance the decline of the forces of natural selection, which 243 indicates that the timing of senescence, in addition to the rate, constitutes a key component of 244 Hamilton's theory of senescence. We thus propose to expand Hamilton's theory to include 245 potential responses of the timing of senescence to increased fertility, by considering both the 246 timing and the intensity of senescence patterns to test reliably evolutionary theories of ageing. 247

248 When looking at sex-differences in senescence patterns, we did not detect any 249 difference in the rate of senescence between male and female wild boar. This finding contrasts 250 with a few studies that reported greater actuarial senescence in males compared to females in 251 large herbivores (Gaillard et al. 2003; Lemaître and Gaillard 2013), a pattern that is generally 252 attributed to the polygynous mating system of these species (Clutton-Brock and Isvaran 2007, 253 Bonduriansky et al. 2008). The red deer (*Cervus elaphus*) case study provides a striking 254 example of higher senescence rates (but not earlier onset) in males than in females (Catchpole 255 et al. 2004). Our results are thus particularly surprising from both empirical (Loison et al.

256 1999, Lemaître and Gaillard 2013) and theoretical (Williams 1957) viewpoints. Indeed, 257 because males had higher mortality rates at all ages from 2 years old onwards than females, 258 they were expected to show greater actuarial senescence (Williams 1957, Hamilton 1966). 259 The higher mortality rates of male wild boar during the whole adult stage might be due to sex-260 specific reproductive tactics and associated mortality risks (Bonduriansky et al. 2008, Festa-261 Bianchet 2013) or to the occurrence of harsh environmental conditions in the study site. 262 Indeed, harsh environments are more likely to influence male than female survival (Toïgo and 263 Gaillard 2003). In wild boar, the differential mortality suffered by males do not translate into stronger senescence. Although the causes of similar senescence rates in both sexes despite 264 265 higher mortality rates in male wild boar remain unknown, it is interesting to notice that 266 similar patterns have been reported in humans (Austad 2011). In both wild boar and humans, 267 the between-sex differences in adult mortality was quite low (about 10%), which might be not 268 enough to trigger substantial between-sex differences in the decline of Hamilton's forces of 269 selection with age. Simulation studies will be required to test this hypothesis.

270

271 Alternatively, differential sources of mortality between sexes could account for the 272 absence of sex differences in actuarial senescence we reported in wild boar. Indeed, recent 273 laboratory experiments have demonstrated that the cause of mortality at early ages can 274 modulate senescence patterns (Chen and Maklakov 2012, Chen et al. 2013). Basically, when 275 the source of mortality is random, all animals in the population have the same probability of 276 dying, independently of their individual phenotype. On the contrary, when mortality is 277 condition-dependent, the probability of dying acts as a filter in the population by removing 278 weakest individuals, which ultimately decreases the intensity of senescence at the population 279 level (Vaupel et al. 1979). In support, an experiment on wild type derived nematodes 280 (Caenorhabditis remanei) has shown high rates of condition-dependent mortality promotes a

deceleration of ageing, independently of sex, within only 12 generations (Chen et al. 2013). In wild boar, differential sources of mortality between sexes could explain our findings in wild boar if condition-dependent mortality only occurs in males. The twofold higher mortality rate of male wild boar at 2 years compared to females might provide the potential filter to keep the best phenotypes at old ages. However, whether mortality at 2 years of age is condition-dependent in males remains to be explored. **ACKNOWLEDGMENTS** We thank the Segretariato per la Presidenza della Repubblica for allowing the study at Castelporziano, for its financial support and the data availability. LITERATURE CITED 

- Austad, S.N. 2011. Handbook of the biology of aging. Academic press, Ed: Masoro, E.J., andS.N. Austad.
- 306 Bininda-Emonds, O.R.P., M. Cardillo, K.E. Jones, R.D.E. MacPhee, R.M.D. Beck, R.
- Grenyer, S.A. Price, R.A. Vos, J.L. Gittleman, and A. Purvis. 2007. The delayed rise of
   present-day mammals. Nature 446:507–512.
- 309 Bonduriansky, R., A. Maklakov, F. Zajitschek, and R. Brooks. 2008. Sexual selection, sexual
- 310 conflict and the evolution of ageing and life span. Funct. Ecol. 22:443-453.
- 311 Burnham, K.P., and D.R. Anderson. 2002. Model selection and multimodel inference: a
- 312 practical information-theoretic approach. 2<sup>nd</sup> edition. Springer Verlag, New-York,
- 313 USA.
- Calder, W.A. 1984. Size, function, and life history. Harvard University Press, Cambridge,
  MA.
- 316 Carey, J.R., and D.S. Judge. 2000. Longevity records: life spans of mammals, birds,

317 amphibians, reptiles and fish. Odense University Press, Odense.

- 318 Carranza, J. 1996. Sexual selection for male body mass and the evolution of litter size in
- 319 mammals. Am. Nat. 148:81-100.
- 320 Catchpole, Y.F., B.J.T. Morgan, T.H. Clutton-Brock, and T. Coulson. 2004. Dimorphism,
- 321 survival and dispersal in red deer. JABES 9:1-26.
- 322 Chen, H., and A.A. Maklakov. 2012. Longer life span evolves under high rates of condition323 dependent mortality. Current Biology 22:2140-2143.
- Chen, H., F. Zajitschek, and A.A. Maklakov. 2013. Why ageing stops: heterogeneity explains
  late-life mortality deceleration in nematodes. Biol. Lett. 9:20130217.
- 326 Choquet, R., J.D. Lebreton, O. Gimenez, A.M. Reboulet, and R. Pradel. 2009a. U-CARE:
- 327 Utilities for performing goodness of fit tests and manipulating CApture-REcapture
- 328 data. Ecography 32:1071-1074.

329	Choquet, R., L. Rouan, and R. Pradel. 2009b. Program E-Surge: a software application for
330	fitting multievent models. Pages 845-865 in D. L. Thomson, E. G. Cooch, and M. J.
331	Conroy, eds. Modeling Demographic Processes In Marked Populations. Volume 3.
332	Springer-Verlag, New York.
333	Clutton-Brock, T.H., and B.C. Sheldon. 2010. Individuals and populations: the role of long-
334	term, individual-based studies of animals in ecology and evolutionary biology. TREE
335	25:562-573.
336	Clutton-Brock, T. H., and K. Isvaran. 2007. Sex differences in ageing in natural populations of
337	vertebrates. Proc. R. Soc. B. 274:3097-3104.
338	Delgado, R., P. Fernandez-Llario, M. Azevedo, A. Beja-Pereira, and P. Santos. 2008. Paternity
339	assessment in free-ranging wild boar (Sus scrofa) – Are littermates full-sibs? Mamm.
340	Biol. 73:169-176.
341	Festa-Bianchet, M. 2013. Human-induced reversal of fortunes for migratory ungulates?
342	Ecology 94:1243-1244.

343 Focardi, S., J.M. Gaillard, F. Ronchi, and S. Rossi, S. 2008. Survival of wild boars in a

- variable environment: unexpected life-history variation in an unusual ungulate. J.
  Mamm. 89:1113-1123.
- 346 Franzetti, B., F. Ronchi, F. Marini, M. Scacco, R. Calmanti, A. Calabrese, A. Paola, M. Paolo,

and S. Focardi. 2012. Nocturnal line transect sampling of wild boar (Sus scrofa) in a

- 348 Mediterranean forest: long-term comparison with capture-mark-resight population
- 349 estimates. Eur. J. Wild. Res. 58:385-402.
- Freckleton, R., P. Harvey, and M. Pagel. 2002. Phylogenetic analysis and comparative data: A
  test and review of evidence. Am. Nat. 160:712–726.
- 352 Gaillard, J.M., D. Pontier, S. Brandt, J.M. Jullien, and D. Allaine. 1992. Sex differentiation in
- 353 postnatal growth rate: a test in a wild boar population. Oecologia 90:167-171.

- 354 Gaillard, J.M., A. Loison, M. Festa-Bianchet, N.G. Yoccoz, and E. Solberg. 2003. Ecological
- 355 correlates of lifespan in populations of large herbivorous mammals. From: Carey,
- 356 James R. and Shripad Tuljapurkar (eds.) Life Span: Evolutionary, Ecological, and
- 357 *Demographic Perspectives*, Supplement to *Population and Development Review*, 29:
- 358 2003. New York: Population Council.
- 359 Gaillard, J.M., N.G. Yoccoz, J.D. Lebreton, C. Bonenfant, S. Devillard, A. Loison, D. Pontier,
- and D. Allainé. 2005. Generation time: a reliable metric to measure life-history

361 variation among mammalian populations. Am. Nat. 166:119-123.

- 362 Gamelon, M., J.M. Gaillard, E. Baubet, S. Devillard, L. Say, S. Brandt, and O. Gimenez.
- 363 2013. The relationship between phenotypic variation among offspring and mother
  364 body mass in wild boar: Evidence of coin-flipping? J. Anim. Ecol. 82:937-945.
- Gimenez O., R. Choquet, and J.D. Lebreton. 2003. Parameter redundancy in multistate
  capture-recapture models. Biometrical J. 45:704-722.
- 367 Hamilton, W.D. 1966. The moulding of senescence by natural selection. J. Theor. Biol. 12:12-368 45.
- 369 Hammers, M., D.S. Richardson, T. Burke, and J. Komdeur. 2013. The impact of reproductive
- 370 investment and early-life environmental conditions on senescence: support for the

disposable soma hypothesis. J. Evol. Biol. 26:1999-2007.

- Harvey, P.H., and M.D. Pagel. 1991. The Comparative Method in Evolutionary Biology. 1<sup>st</sup>
  ed. Oxford University Press, USA.
- 374 Jones, O.R., J.M. Gaillard, S. Tuljapurkar, J.S. Alho, K.B. Armitage, P.H. Becker et al. 2008.
- 375 Senescence rates are determined by ranking on the fast–slow life-history continuum.
- 376 Ecol. Lett. 11:664-673.
- 377 Kirkwood, T.B.L. 1977. Evolution of ageing. Nature 270:301-304.
- 378 Kirkwood, T.B.L. and S.N. Austad. 2000. Why do we age? Nature 408:233-238.

- Lebreton, J.-D., K.P. Burnham, J. Clobert, and D.R. Anderson. 1992. Modeling survival and
  testing biological hypotheses using marked animals: a unified approach with case
  studies. Ecol. Monogr. 62:67-118.
- Lemaître, J.F. and J.M. Gaillard. 2013. Male survival patterns do not depend on male
  allocation to sexual competition in large herbivores. Behav. Ecol. 24:421-428.
- 384 Loison, A., M. Festa-Bianchet, J.M. Gaillard, J.T. Jorgenson, and J.M. Jullien. 1999. Age-
- 385 specific survival in five populations of ungulates: evidence of senescence. Ecology
  386 80:2539-2554.
- Maklakov, A.A., and V. Lummaa. 2013. Evolution of sex differences in lifespan and aging:
  causes and constraints. Bioessays 35:717-724.
- 389 Medawar, P. B. 1952. An Unsolved Problem of Biology (Lewis, London 1952).
- Nichols, J.D., J.E. Hines, and P. Blums. 1997. Tests for senescent decline in annual survival
  probabilities of common pochards, *Aythya ferina*. Ecology 78:1009-1018.
- 392 Nussey, D.H., L.E.B. Kruuk, A. Donald, M. Fowlie, and T.H. Clutton-Brock. 2006. The rate
- 393 of senescence in maternal performance increases with early-life fecundity in red deer.
  394 Ecol. Lett. 9:1342-1350.
- 395 Nussey, D.H., H. Froy, J.F. Lemaître, J.M. Gaillard, and S.N. Austad. 2013. Senescence in
- 396 natural populations of animals: Widespread evidence and its implications for bio-
- 397 gerontology. Ageing Res. Rev. 12:214-225.
- 398 Péron, G., O. Gimenez, A. Charmantier, J.M. Gaillard and P.A. Crochet. 2010. Age at onset of
- senescence in birds and mammals is predicted by early-life performance. Proc. R. Soc.
  B. 277:2849-2856.
- 401 Pradel, R., O. Gimenez, and J.D. Lebreton. 2005. Principles and interest of GOF tests for
  402 multistate capture-recapture models. Anim. Biodivers. Conserv. 28:189-204.

- 403 Pradel, R. and A. Sanz-Aguilar. 2012. Modeling trap-awareness and related phenomena in
  404 capture-recapture studies. PlosOne. 7:e32666.
- 405 R Development Core Team. 2012. R: A Language and Environment for Statistical

406 Computing. R Foundation for Statistical Computing, Vienna, Austria.

- 407 Revell, L.J. 2010. Phylogenetic signal and linear regression on species data. Methods in Ecol.
  408 Evol. 1:319-329.
- 409 Ricklefs, R.E. 2010. Life-history connections to rates of aging in terrestrial vertebrates. PNAS
  410 107:10314-10319.

411 Servanty, S., J.M. Gaillard, D. Allainé, S. Brandt, and E. Baubet. 2007. Litter size and fetal
412 sex ratio adjustment in a highly polytocous species: the wild boar. Behav. Ecol.
413 18:427-432.

414 Servanty, S., J.M. Gaillard, C. Toïgo, S. Brandt, and E. Baubet. 2009. Pulsed resources and
415 climate-induced variation in the reproductive traits of wild boar under high hunting
416 pressure. J. Anim. Ecol. 78:1278-1290.

417 Toïgo, C. and J.M. Gaillard. 2003. Causes of sex-biased adult survival in ungulates: sexual
418 size dimorphism, mating tactic or environment harshness? Oikos 101:376-384.

419 Vaupel, J.W., K.G. Manton, and E. Stallard. Impact of heterogeneity in individual frailty on

420 the dynamics of mortality. Demography 16:439-454.

- 421 Vinogradov, A.E. 1998. Male reproductive strategy and decrease longevity. Acta Biotheor.
- 422 46:157-160.
- Williams, G.C. 1957. Pleiotropy, natural selection, and the evolution of senescence. Evolution
  11:398-411.
- 425
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	Model notation Np AIC AAIC
450	the 10 best models and the difference between each tested model and the best one ( $\Delta AIC$ ).
449	age and survival from x year of age. Displayed are the number of parameters (Np), the AIC of
448	Castelporziano, Italy. $\Phi_{a(x-13)}$ corresponds to a linear relationship (on a logit-scale) between
447	Table 1. Model selection for age-specific survival patterns in the wild boar population of
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Model notation	Np	AIC	ΔΑΙC
	-		

$\Phi_1[\Phi_2]$ .Sex $[\Phi_{a(3-13)}]$ +Sex	25	6422.88	0.00	
$[\Phi_1\Phi_2]$ .Sex $[\Phi_{a(3-13)}]$ +Sex	26	6423.06	0.18	
$\Phi_1[\Phi_2 \ \Phi_{a(3-13)}]$ .Sex	26	6424.74	1.86	
$[\Phi_1\Phi_2\Phi_{a(3-13)}].Sex$	27	6424.94	2.06	
$\Phi_1[\Phi_2].Sex \ \Phi_3[\Phi_{a(4-13)}]+Sex$	26	6425.24	2.36	
$[\Phi_1\Phi_2]$ .Sex $\Phi_3[\Phi_{a(4-13)}]$ +Sex	27	6425.35	2.47	
$\Phi_1[\Phi_2]$ .Sex $\Phi_3[\Phi_{a(4-13)}]$ .Sex	27	6426.61	3.73	
$\Phi_1[\Phi_2\Phi_3]$ .Sex $[\Phi_{a(4-13)}]$ +Sex	27	6426.81	3.93	
$[\Phi_1\Phi_2\Phi_3]$ .Sex $[\Phi_{a(4-13)}]$ +Sex	28	6427.46	4.58	
$\Phi_1[\Phi_2\Phi_3\Phi_{a(4-13)}]$ .Sex	28	6428.38	5.50	

### 452 Figure legends

453 Figure 1. Age-specific survival for the wild boar population of Castelporziano, Italy (grey for 454 females, black for males). Filled circles display the estimates from the full age-dependent 455 model (with standard errors). The line corresponds to the selected model involving a linear 456 decrease (on a logit-scale) of survival from 3 years old onwards. Numbers denote wild boars 457 captured for last time in each age class.

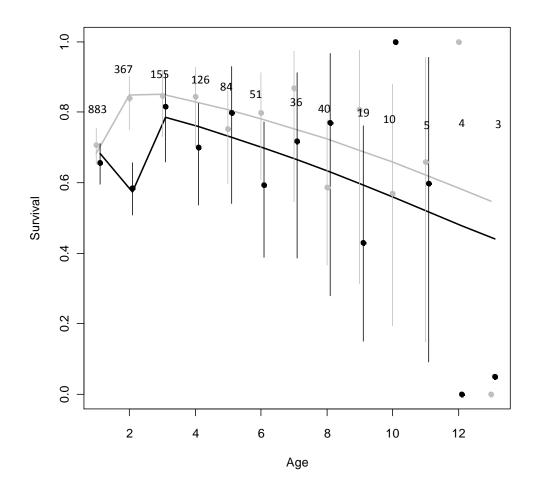
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459 Figure 2. Relationship between log-transformed mortality rates and age in the wild boar
460 population of Castelporziano, Italy (grey for females, black for males).

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Figure 3. Relationships (a) between senescence rates and body mass (in g) (on a log-log 462 463 scale) and (b) between onsets of senescence (in years) and body mass (in g) (on a log-log 464 scale) for females in populations belonging to 7 different ungulate species (values obtained from Gaillard et al. 2003 and Péron et al. 2010) for 1- Antilocapra americana at Montana 465 466 (USA), 2- Capra ibex, at Belledone (France), 3- Capreolus capreolus at Trois-Fontaines 467 (France), 4- Capreolus capreolus at Chizé (France), 5- Oreamnos americanus at Caw Ridge 468 (Canada), 6- Ovis canadensis at Sheep River (Canada), 7- Ovis canadensis at Ram Mountain 469 (Canada), 8- Cervus elaphus at Rum (Scotland), 9- Rupicapra pyrenaica at Orlu (France). The 470 observed values for wild boar females (from this study, red circle, not included in the 471 allometic relationship) and the expected values for wild boar from the allometric relationship 472 (red square) are displayed. The comparison shows that wild boar females have a rate of 473 senescence close to the expectation, but a much earlier onset of senescence.

**Figure 1.** 



**Figure 2.** 

